

## EXAM C QUESTION OF THE WEEK

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### Week of April 14/08

The prior distribution of the parameter  $\lambda$  is a gamma distribution with parameters  $\alpha = 2$  and  $\theta = 2$ . The model distribution of  $X$  given  $\lambda$  is Poisson with a mean of  $\lambda$ . A sample value of  $X$  is observed, say  $X_1$ , based on a random value of  $\lambda$  from the prior distribution. Another sample value is then obtained, say  $X_2$ , based on the same (unknown) value of  $\lambda$ . Find the probability that  $X_2$  is at least 1 given that  $X_1$  is at least 1.

**The solution can be found below.**

## Week of April 14/08 - Solution

The unconditional distribution of  $X_1$  is negative binomial with  $r = \alpha$  and  $\beta = \theta$ .

$$P[X_2 \geq 1 | X_1 \geq 1] = \frac{P[(X_2 \geq 1) \cap (X_1 \geq 1)]}{P[X_1 \geq 1]} .$$

$$P[X_1 \geq 1] = 1 - P[X_1 = 0] = 1 - \left(\frac{1}{1+\beta}\right)^r = 1 - \left(\frac{1}{1+\theta}\right)^\alpha = 1 - \left(\frac{1}{3}\right)^2 = \frac{8}{9} .$$

$$\begin{aligned} P[(X_2 \geq 1) \cap (X_1 \geq 1)] &= \int_0^\infty P[(X_2 \geq 1) \cap (X_1 \geq 1) | \lambda] \cdot \pi(\lambda) d\lambda \\ &= \int_0^\infty P[X_2 \geq 1] \cdot P[X_1 \geq 1 | \lambda] \cdot \pi(\lambda) d\lambda \\ &= \int_0^\infty (1 - e^{-\lambda})(1 - e^{-\lambda}) \frac{\lambda^{\alpha-1} e^{-\lambda/\theta}}{\theta^\alpha \Gamma(\alpha)} d\lambda = \int_0^\infty (1 - e^{-\lambda})(1 - e^{-\lambda}) \frac{\lambda^1 e^{-\lambda/2}}{4 \Gamma(2)} d\lambda \\ &= \frac{1}{4} \cdot \int_0^\infty \lambda (e^{-\lambda/2} - 2e^{-\lambda(\frac{3}{2})} + e^{-\lambda(\frac{5}{2})}) d\lambda \\ &= \frac{1}{4} \cdot \left[ \frac{1}{(\frac{1}{2})^2} - 2 \frac{1}{(\frac{3}{2})^2} + \frac{1}{(\frac{5}{2})^2} \right] = 1 - \frac{2}{3^2} + \frac{1}{5^2} = \frac{184}{225} . \end{aligned}$$

$$P[X_2 \geq 1 | X_1 \geq 1] = \frac{184/225}{8/9} = .92 .$$